



COS 484/584

Expectation Maximization

Spring 2021

(some slides adapted from Regina Barzilay and Michael Collins)

Midterm

- Logistics announced on Canvas
 - March 10, 12pm ET - March 11, 12pm ET
- Please fill out the survey on your preferred time for taking the exam so we can better plan email support
- Midterm review: COS 484 precept this week (March 5)
- TAs have posted a survey on Canvas - please fill it out if you'd like them to review specific topics

Expectation Maximization

- If we have **partially observable data**, x_i examples only, then

$$L(\theta) = \sum_i \log \sum_{y \in \mathcal{Y}} P(x_i, y | \theta)$$

- The EM (Expectation Maximization) algorithm is a method for finding

$$\theta_{MLE} = \arg \max_{\theta} L(\theta) = \arg \max_{\theta} \sum_i \log \sum_{y \in \mathcal{Y}} P(x_i, y | \theta)$$

The three coins example

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- In the three coins example,

$\mathcal{Y} = \{H, T\}$ (possible outcomes of coin 0)

$\mathcal{X} = \{HHH, TTT, HTT, THH, HHT, TTH, HTH, THT\}$

$\theta = \{\lambda, p_1, p_2\}$

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(all possible
observations of length 3)

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(all possible
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$\theta = \{\lambda, p_1, p_2\}$

- and $P(x, y | \theta) = P(y | \theta) P(x | y, \theta)$

where

$$P(y | \theta) = \begin{cases} \lambda & \text{if } y = H \\ 1 - \lambda & \text{if } y = T \end{cases}$$

and

$$P(x | y, \theta) = \begin{cases} p_1^h (1 - p_1)^t & \text{if } y = H \\ p_2^h (1 - p_2)^t & \text{if } y = T \end{cases}$$

The three coins example

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- Calculating various probabilities:

$$P(x = THT, y = H | \theta) = \lambda p_1 (1 - p_1)^2$$

$$P(x = THT, y = T | \theta) = (1 - \lambda) p_2 (1 - p_2)^2$$

The three coins example

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$$P(x = THT, y = H | \theta) = \lambda p_1 (1 - p_1)^2$$

$$P(x = THT, y = T | \theta) = (1 - \lambda) p_2 (1 - p_2)^2$$

$$\begin{aligned} P(x = THT | \theta) &= P(x = THT, y = H | \theta) + P(x = THT, y = T | \theta) \\ &= \lambda p_1 (1 - p_1)^2 + (1 - \lambda) p_2 (1 - p_2)^2 \end{aligned}$$

$$\begin{aligned} P(y = H | x = THT, \theta) &= \frac{P(x = THT, y = H | \theta)}{P(x = THT | \theta)} \\ &= \frac{\lambda p_1 (1 - p_1)^2}{\lambda p_1 (1 - p_1)^2 + (1 - \lambda) p_2 (1 - p_2)^2} \end{aligned}$$

The three coins example

- New estimates:

$(\langle \text{HHH} \rangle, H)$	$P(y = H \mid \text{HHH}) = 0.0508$
$(\langle \text{HHH} \rangle, T)$	$P(y = T \mid \text{HHH}) = 0.9492$
$(\langle \text{TTT} \rangle, H)$	$P(y = H \mid \text{TTT}) = 0.6967$
$(\langle \text{TTT} \rangle, T)$	$P(y = T \mid \text{TTT}) = 0.3033$
$(\langle \text{HHH} \rangle, H)$	$P(y = H \mid \text{HHH}) = 0.0508$
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$(\langle \text{HHH} \rangle, T)$	$P(y = T \mid \text{HHH}) = 0.9492$

$$\lambda = \frac{3 \times 0.0508 + 2 \times 0.6967}{5} = 0.3092$$

$$p_1 = \frac{3 \times 3 \times 0.0508 + 0 \times 2 \times 0.6967}{3 \times 3 \times 0.0508 + 3 \times 2 \times 0.6967} = 0.0987$$

$$p_2 = \frac{3 \times 3 \times 0.9492 + 0 \times 2 \times 0.3033}{3 \times 3 \times 0.9492 + 3 \times 2 \times 0.3033} = 0.8244$$

The three coins example

- New estimates:

$$(\langle \text{HHH} \rangle, H) \quad P(y = H \mid \text{HHH}) = 0.0508$$

$$(\langle \text{HHH} \rangle, T) \quad P(y = T \mid \text{HHH}) = 0.9492$$

$$(\langle \text{TTT} \rangle, H) \quad P(y = H \mid \text{TTT}) = 0.6967$$

$$(\langle \text{TTT} \rangle, T) \quad P(y = T \mid \text{TTT}) = 0.3033$$

$$(\langle \text{HHH} \rangle, H) \quad P(y = H \mid \text{HHH}) = 0.0508$$

$$(\langle \text{HHH} \rangle, T) \quad P(y = T \mid \text{HHH}) = 0.9492$$

$$(\langle \text{TTT} \rangle, H) \quad P(y = H \mid \text{TTT}) = 0.6967$$

$$(\langle \text{TTT} \rangle, T) \quad P(y = T \mid \text{TTT}) = 0.3033$$

$$(\langle \text{HHH} \rangle, H) \quad P(y = H \mid \text{HHH}) = 0.0508$$

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$$\lambda = \frac{3 \times 0.0508 + 2 \times 0.6967}{5} = 0.3092$$

$$p_1 = \frac{3 \times 3 \times 0.0508 + 0 \times 2 \times 0.6967}{3 \times 3 \times 0.0508 + 3 \times 2 \times 0.6967} = 0.0987$$

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- Begin with parameters: $\lambda = 0.3$, $p_1 = 0.3$, $p_2 = 0.6$

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$$P(y = H | x = \langle HHH \rangle) = 0.0508$$

$$P(y = H | x = \langle TTT \rangle) = 0.6967$$

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$$P(y = H | x = \langle HHH \rangle) = 0.0508$$
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- Re-estimate parameters to be
$$\lambda = 0.3092, p_1 = 0.0987, p_2 = 0.8244$$

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- Re-estimate parameters to be
$$\lambda = 0.3092, p_1 = 0.0987, p_2 = 0.8244$$

Repeat!



EM iterations (example I)

$$P(y = H | x_1)$$

Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
0	0.3000	0.3000	0.6000	0.0508	0.6967	0.0508	0.6967
1	0.3738	0.0680	0.7578	0.0004	0.9714	0.0004	0.9714
2	0.4859	0.0004	0.9722	0.0000	1.0000	0.0000	1.0000
3	0.5000	0.0000	1.0000	0.0000	1.0000	0.0000	1.0000

The coin example for $x = \{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$. The solution that EM reaches is intuitively correct: the coin tosser has two coins, one which always shows heads, and another which always shows tails, and is picking between them with equal probability ($\lambda = 0.5$) using coin 0.

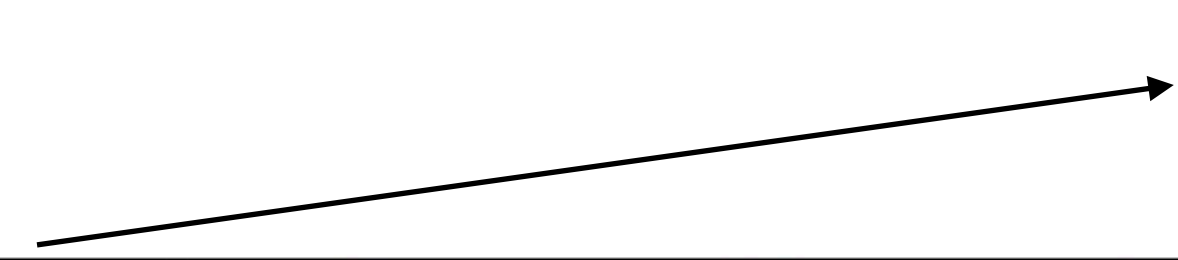
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The coin example for $x = \{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$. The solution that EM reaches is intuitively correct: the coin tosser has two coins, one which always shows heads, and another which always shows tails, and is picking between them with equal probability ($\lambda = 0.5$) using coin 0.

Posterior probabilities \bar{p}_i show that we are certain that coin 1 (tail-biased) generated x_2 and x_4 , whereas coin 2 generated x_1 and x_3

EM iterations (example 2)

$$P(y = H | x_1)$$

Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4	\bar{p}_5
0	0.3000	0.3000	0.6000	0.0508	0.6967	0.0508	0.6967	0.0508

Coin example for $\{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle\}$

EM iterations (example 2)



$$P(y = H | x_1)$$

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Coin example for $\{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle\}$

Which of these would you expect EM to converge to?

A) $\lambda = 0.5, p_1 = 0.5, p_2 = 0.5$

B) $\lambda = 0.5, p_1 = 1, p_2 = 0$

C) $\lambda = 0.4, p_1 = 0, p_2 = 1$

EM iterations (example 2)



$$P(y = H | x_1)$$

Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4	\bar{p}_5
0	0.3000	0.3000	0.6000	0.0508	0.6967	0.0508	0.6967	0.0508
1	0.3092	0.0987	0.8244	0.0008	0.9837	0.0008	0.9837	0.0008
2	0.3940	0.0012	0.9893	0.0000	1.0000	0.0000	1.0000	0.0000
3	0.4000	0.0000	1.0000	0.0000	1.0000	0.0000	1.0000	0.0000

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Coin example for $\{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle\}$

Which of these would you expect EM to converge to?

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B) $\lambda = 0.5, p_1 = 1, p_2 = 0$

C) $\lambda = 0.4, p_1 = 0, p_2 = 1$

λ is now 0.4, indicating that coin 0 has a probability 0.4 of selecting the tail-biased coin 1

EM iterations (example 3)

Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
0	0.3000	0.3000	0.6000	0.1579	0.6967	0.0508	0.6967

Coin example for $x = \{\langle HHT \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

EM iterations (example 3)



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Coin example for $x = \{\langle HHT \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

Which of these would you expect EM to converge to?

A) $\lambda = 0.49$, $p_1 = 0.12$, $p_2 = 0$

B) $\lambda = 0.49$, $p_1 = 0$, $p_2 = 0.82$

C) $\lambda = 0.5$, $p_1 = 0.5$, $p_2 = 0.5$

EM iterations (example 3)



Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
0	0.3000	0.3000	0.6000	0.1579	0.6967	0.0508	0.6967
1	0.4005	0.0974	0.6300	0.0375	0.9065	0.0025	0.9065
2	0.4632	0.0148	0.7635	0.0014	0.9842	0.0000	0.9842
3	0.4924	0.0005	0.8205	0.0000	0.9941	0.0000	0.9941
4	0.4970	0.0000	0.8284	0.0000	0.9949	0.0000	0.9949

Coin example for $x = \{\langle HHT \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

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Coin example for $x = \{\langle HHT \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

- EM selects a tails-only coin ($p_1 = 0$), and a coin which is heavily heads-biased ($p_2 = 0.8284$).
- It is certain that x_1 and x_3 were generated by coin 2 since they contain heads.
- x_2 and x_4 could have been generated by either coin but coin 1 (tail-biased) is far more likely.

EM iterations (example 4)



Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
0	0.3000	0.7000	0.7000	0.3000	0.3000	0.3000	0.3000

Coin example for $x = \{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

EM iterations (example 4)



Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
0	0.3000	0.7000	0.7000	0.3000	0.3000	0.3000	0.3000

Coin example for $x = \{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

Which of these would you expect EM to converge to?

A) $\lambda = 0.3, p_1 = 0.5, p_2 = 0.5$

B) $\lambda = 0.5, p_1 = 0.5, p_2 = 0.5$

C) $\lambda = 0.5, p_1 = 0, p_2 = 1$

EM iterations (example 4)



Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
0	0.3000	0.7000	0.7000	0.3000	0.3000	0.3000	0.3000

Coin example for $x = \{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

Which of these would you expect EM to converge to?

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EM iterations (example 4)



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2	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
3	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
4	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
5	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
6	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000

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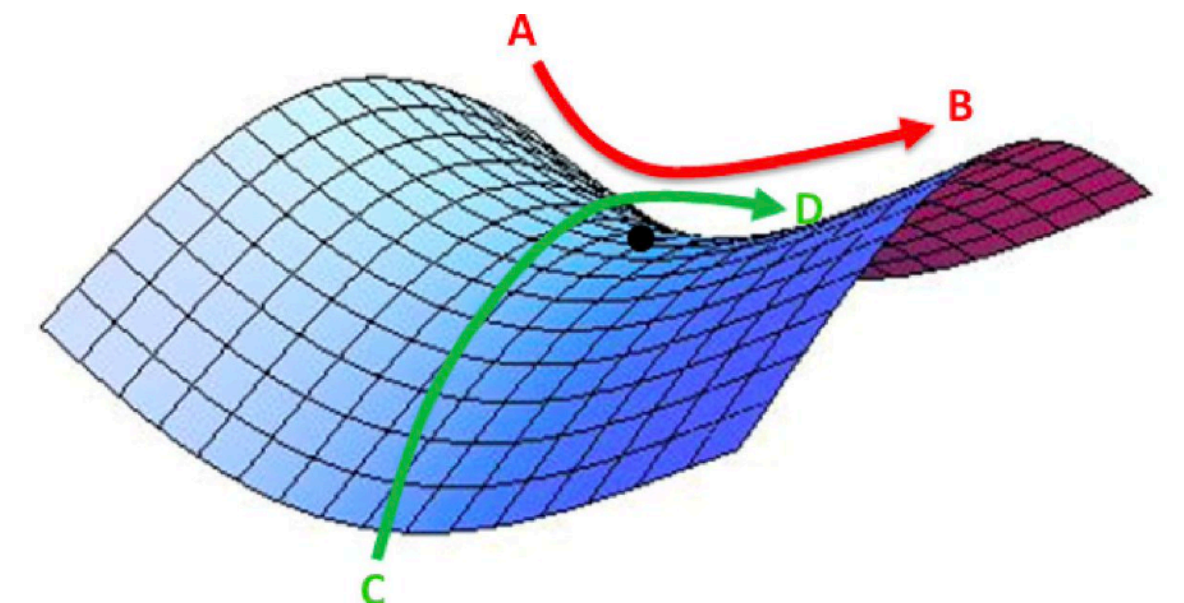
Initialization matters



Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
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1	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
2	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
3	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
4	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
5	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000
6	0.3000	0.5000	0.5000	0.3000	0.3000	0.3000	0.3000

Coin example for $x = \{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

In this case, EM is stuck at a **saddle point**.



Iteration	λ	p_1	p_2	\bar{p}_1	\bar{p}_2	\bar{p}_3	\bar{p}_4
0	0.3000	0.7001	0.7000	0.3001	0.2998	0.3001	0.2998
1	0.2999	0.5003	0.4999	0.3004	0.2995	0.3004	0.2995
2	0.2999	0.5008	0.4997	0.3013	0.2986	0.3013	0.2986
3	0.2999	0.5023	0.4990	0.3040	0.2959	0.3040	0.2959
4	0.3000	0.5068	0.4971	0.3122	0.2879	0.3122	0.2879
5	0.3000	0.5202	0.4913	0.3373	0.2645	0.3373	0.2645
6	0.3009	0.5605	0.4740	0.4157	0.2007	0.4157	0.2007
7	0.3082	0.6744	0.4223	0.6447	0.0739	0.6447	0.0739
8	0.3593	0.8972	0.2773	0.9500	0.0016	0.9500	0.0016
9	0.4758	0.9983	0.0477	0.9999	0.0000	0.9999	0.0000
10	0.4999	1.0000	0.0001	1.0000	0.0000	1.0000	0.0000
11	0.5000	1.0000	0.0000	1.0000	0.0000	1.0000	0.0000

Coin example for $x = \{\langle HHH \rangle, \langle TTT \rangle, \langle HHH \rangle, \langle TTT \rangle\}$.

If we initialize p_1 and p_2 even a small amount away from the saddle point $p_1 = p_2$, EM diverges and eventually reaches the global maximum

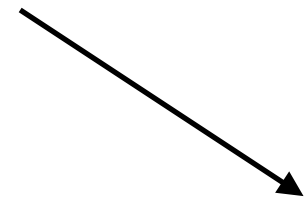
The EM algorithm

The EM algorithm

- θ^t is the parameter vector at the t^{th} iteration

The EM algorithm

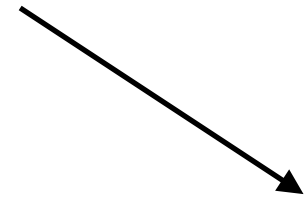
Superscript for iteration #



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$$\theta^t = \arg \max_{\theta} Q(\theta, \theta^{t-1})$$

where

$$Q(\theta, \theta^{t-1}) = \sum_i \sum_{y \in \mathcal{Y}} P(y | x_i, \theta^{t-1}) \log P(x_i, y | \theta)$$

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How did we get $\arg \max_{\theta} Q$ from $\arg \max_{\theta} \sum_i \log \sum_{y \in \mathcal{Y}} P(x_i, y | \theta)$? \Rightarrow Jensen's inequality
(advanced; see optional reading from Andrew Ng)

The EM algorithm

- θ^t is the parameter vector at the t^{th} iteration
- Choose θ^0 at random (or using smart heuristics)
- (E step): Compute *expected* counts for every parameter θ_r :

$$\overline{Count}(r) = \sum_{i=1}^n \sum_y P(y | x_i, \theta^{t-1}) Count(x_i, y, r)$$

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- (M step): Re-estimate parameters using expected counts to **maximize likelihood (MLE estimate)**

$$\text{e.g. } \theta_{DT \rightarrow NN} = \frac{\overline{Count}(DT \rightarrow NN)}{\sum_{\beta} \overline{Count}(DT \rightarrow \beta)}$$

The EM algorithm

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- Key points:
 - Intuition: Fill in hidden variables y according to $P(y | x_i, \theta)$
 - Create a “pseudo-dataset” with fractional counts
 - EM is guaranteed to converge to a **local** maximum, or saddle-point, of the likelihood function
- In general, if $\arg \max_{\theta} \sum_i \log P(x_i, y_i | \theta)$ has a simple analytic solution, then
$$\arg \max_{\theta} \sum_i \sum_y P(y | x_i, \theta) \log P(x_i, y | \theta)$$
 also has a simple solution.

Example: EM for HMM

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- We observe only word sequences X_1, X_2, \dots, X_n (no tags Y)

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- Let θ be the vector of all transition parameters (include initial state distribution as a special case, $\emptyset \rightarrow s$)
- Let ϕ be the vector of all emission parameters
- Initialize parameters to some values θ^0 and ϕ^0

Each X and Y is a sequence on its own

Recap: Estimating HMM parameters



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Which of these is the correct MLE estimate for the transition parameter $\theta_{a \rightarrow b}$ of an HMM (where a, b, b' are states) ?

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B) $\theta_{a \rightarrow b} = \frac{\text{Count}(a \rightarrow b)}{\sum_{b'} \text{Count}(a \rightarrow b')}$

C) $\theta_{a \rightarrow b} = \frac{\text{Count}(a \rightarrow b)}{\sum_{a'} \sum_{b'} \text{Count}(a' \rightarrow b')}$

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- Here, counts are estimated by simply checking for occurrence of the transition/emission in every data sequence

$$\text{e.g. } \text{Count}(a \rightarrow b) = \sum_{i=1}^n \text{Count}(X_i, Y_i, a \rightarrow b)$$

(number of times the transition occurs in each data point)

Example: EM for HMM

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- Initialize parameters θ^0 and ϕ^0

Example: EM for HMM

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- **(E-Step)** Compute **expected** counts

$$\overline{Count}(a \rightarrow b) = \sum_{i=1}^n \sum_Y P(Y|X_i, \theta^{t-1}, \phi^{t-1}) Count(X_i, Y, a \rightarrow b)$$

$$= \sum_{i=1}^n \sum_Y P(Y|X_i, \theta^{t-1}, \phi^{t-1}) Count(Y, a \rightarrow b)$$

$$\overline{Count}(a \rightarrow A) = \sum_{i=1}^n \sum_Y P(Y|X_i, \theta^{t-1}, \phi^{t-1}) Count(X_i, Y, a \rightarrow A)$$

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(why?)

$$\overline{Count}(a \rightarrow A) = \sum_{i=1}^n \sum_Y P(Y|X_i, \theta^{t-1}, \phi^{t-1}) Count(X_i, Y, a \rightarrow A)$$

#times transition
appears in Y

previous parameters

Example: EM for HMM

- (M-Step)

$$\theta_{a \rightarrow b}^t = \frac{\overline{Count}(a \rightarrow b)}{\sum_{a \rightarrow b'} \overline{Count}(a \rightarrow b')}$$

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Example: EM for HMM

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Very similar to the MLE update we saw for HMMs

- **(E-Step)** Compute **expected** counts

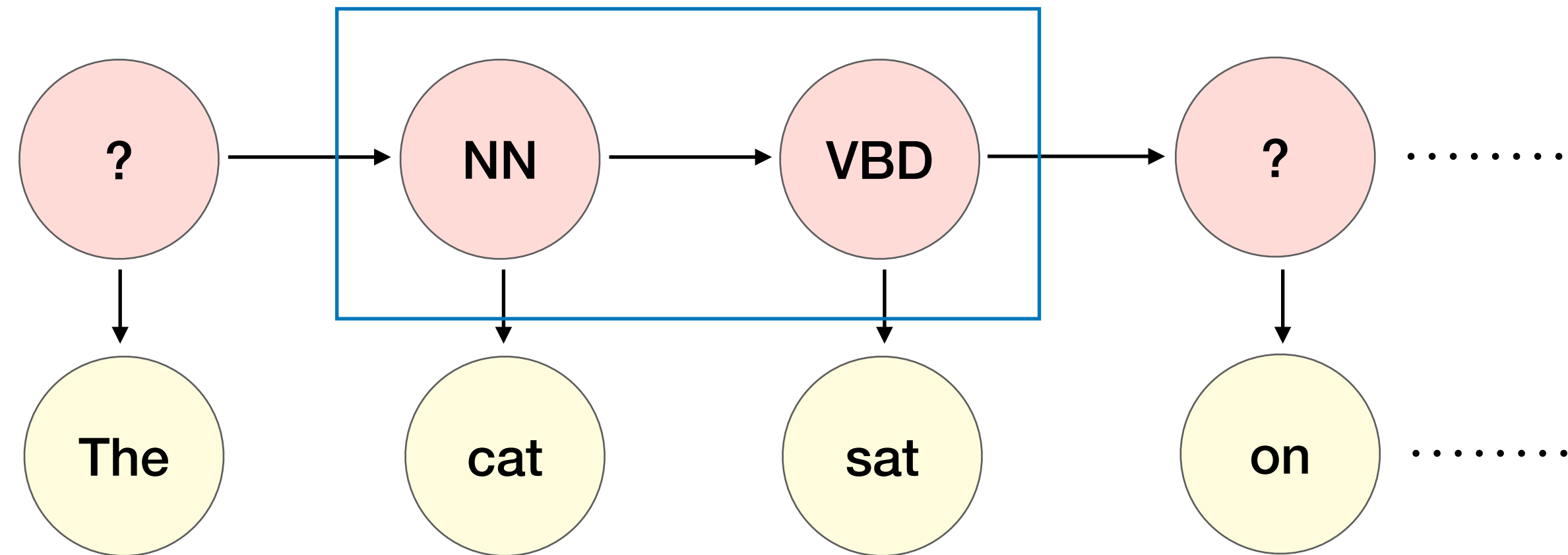
$$\begin{aligned}\overline{Count}(a \rightarrow b) &= \sum_{i=1}^n \sum_Y P(Y|X_i, \theta^{t-1}, \phi^{t-1}) Count(X_i, Y, a \rightarrow b) \\ &= \sum_{i=1}^n \sum_Y P(Y|X_i, \theta^{t-1}, \phi^{t-1}) Count(Y, a \rightarrow b) \\ \overline{Count}(a \rightarrow A) &= \sum_{i=1}^n \sum_Y P(Y|X_i, \theta^{t-1}, \phi^{t-1}) Count(X_i, Y, a \rightarrow A)\end{aligned}$$

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Cannot enumerate all possible Y!

Efficient EM



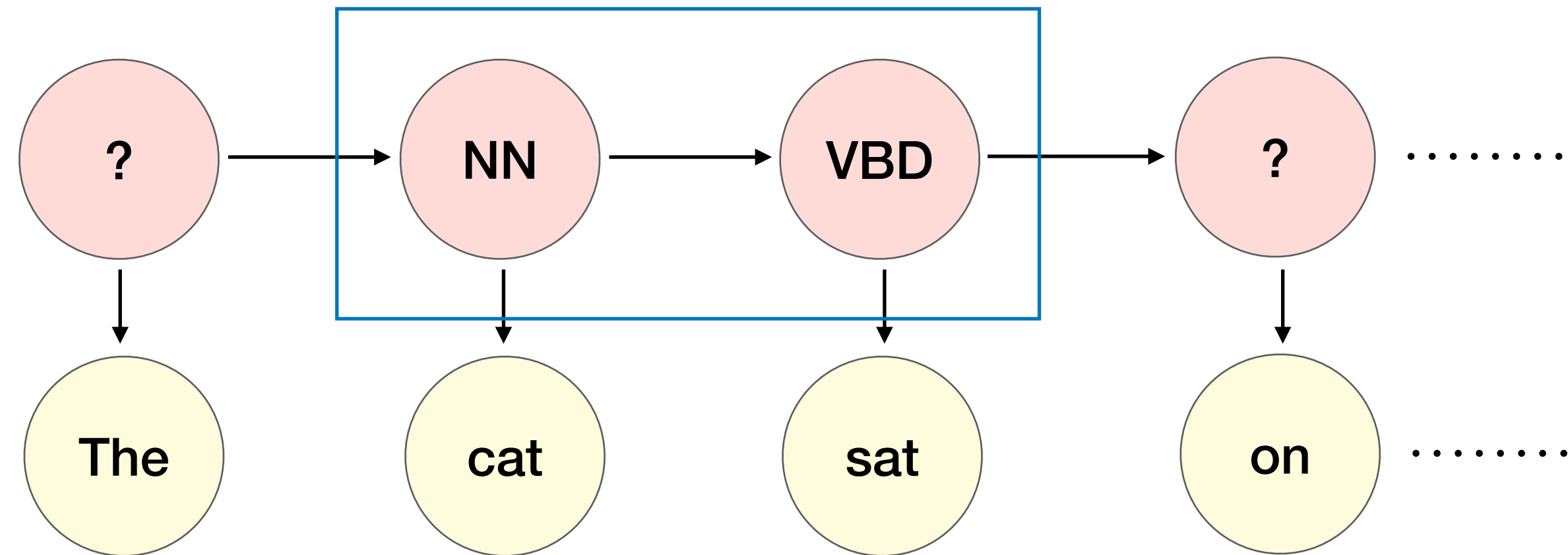
$$Y = \langle y_1, y_2, \dots, y_m \rangle$$

- **(E-Step)**

$$\begin{aligned} \overline{Count}(NN \rightarrow VBD) &= \sum_{i=1}^n \sum_Y P(Y | X_i, \theta^{t-1}, \phi^{t-1}) Count(Y, NN \rightarrow VBD) \\ &= \sum_i \sum_{j=1}^{m-1} P(y_j = NN, y_{j+1} = VBD | X_i, \theta^{t-1}, \phi^{t-1}) \end{aligned}$$

where m is the length of the sequence X_i

Efficient EM



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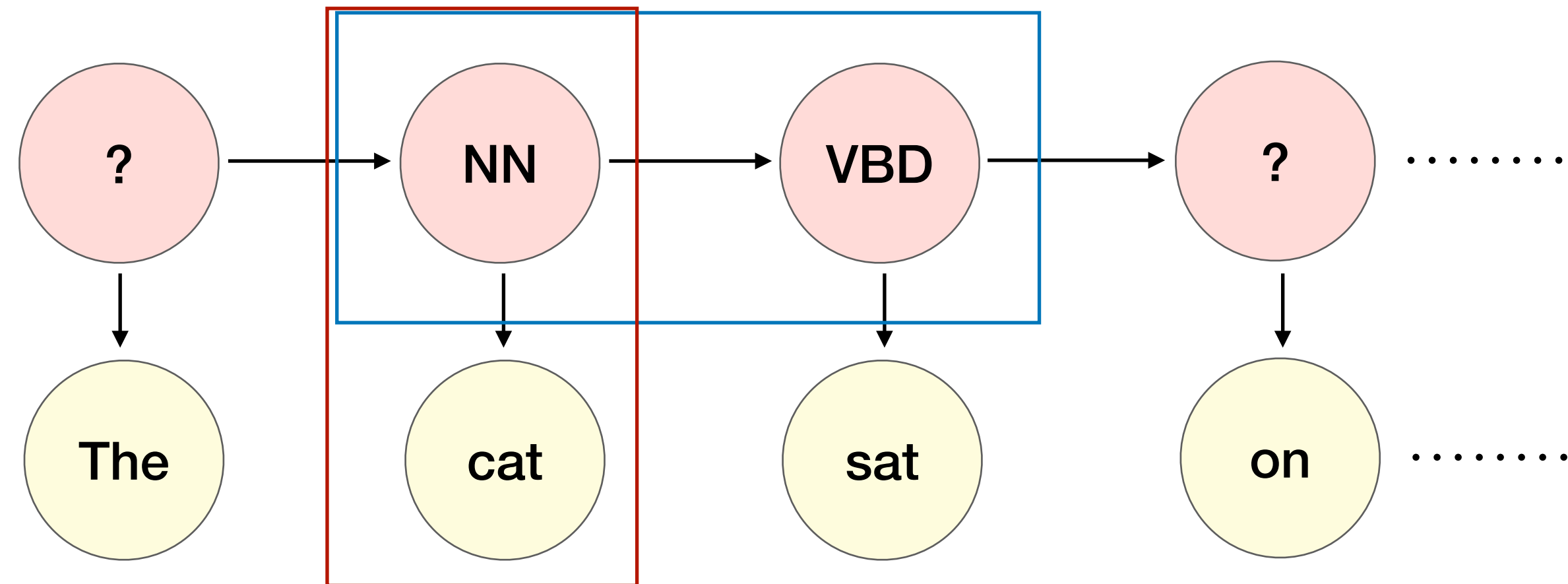
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All other y variables
marginalized out

Efficient EM



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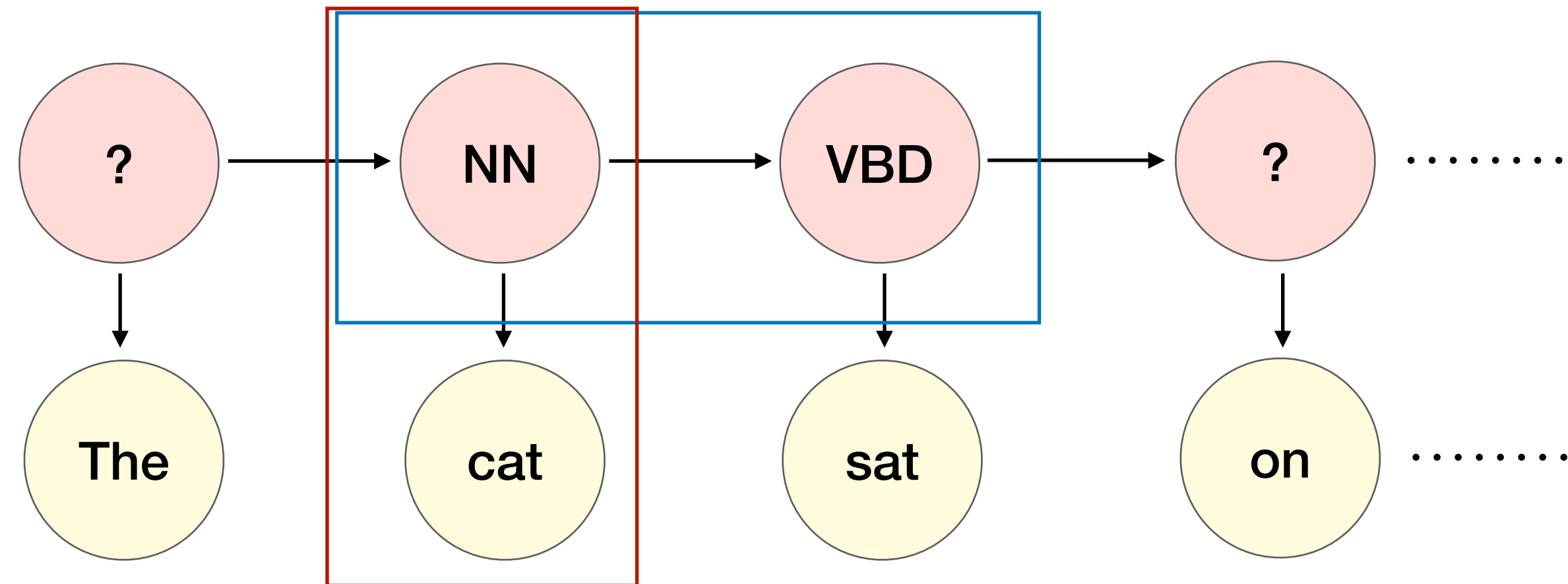
- **(E-Step)**

$$\begin{aligned} \overline{Count}(NN \rightarrow VBD) &= \sum_{i=1}^n \sum_Y P(Y | X_i, \theta^{t-1}, \phi^{t-1}) Count(Y, \theta_k) \\ &= \sum_i \sum_{j=1}^m P(y_j = NN, y_{j+1} = VBD | X_i, \theta^{t-1}, \phi^{t-1}) \end{aligned}$$

where m is the length of the sequence X_i

$$\text{Similarly, } \overline{Count}(NN \rightarrow cat) = \sum_i \sum_{j: X_{ij} = cat} P(y_j = NN | X_i, \theta^{t-1}, \phi^{t-1})$$

Efficient EM



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- **(E-Step)**

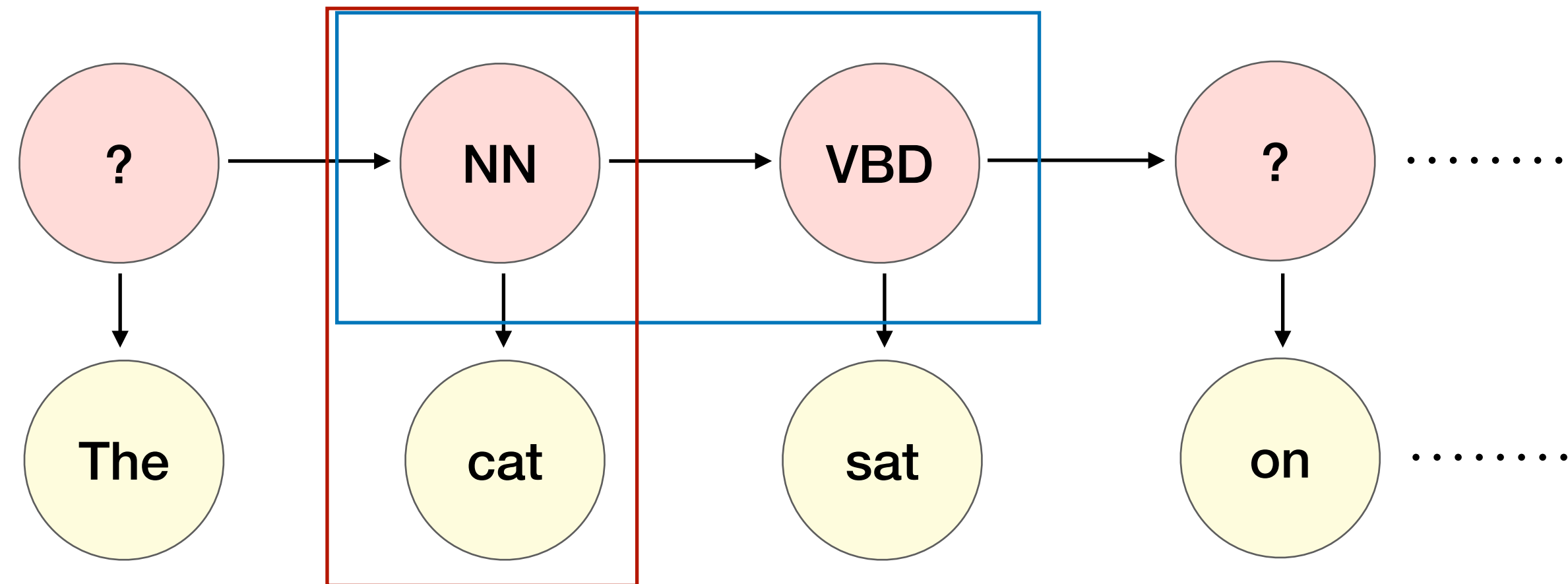
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Efficient EM



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only indices where the
word is 'cat'

Forward-backward algorithm

Forward-backward algorithm

- Define:

$$\alpha_s(j) = P(x_1, \dots, x_{j-1}, y_j = s \mid \theta, \phi) \quad \text{(forward probability)}$$

i.e. the marginal probability of seeing observations x_1, \dots, x_{j-1} and the particular state s in the j^{th} position.

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- $\beta_s(j) = P(x_j, \dots, x_m \mid y_j = s, \theta, \phi) \quad \text{(backward probability)}$

i.e. the marginal probability of seeing observations x_j, \dots, x_m given $y_j = s$

Forward-backward algorithm

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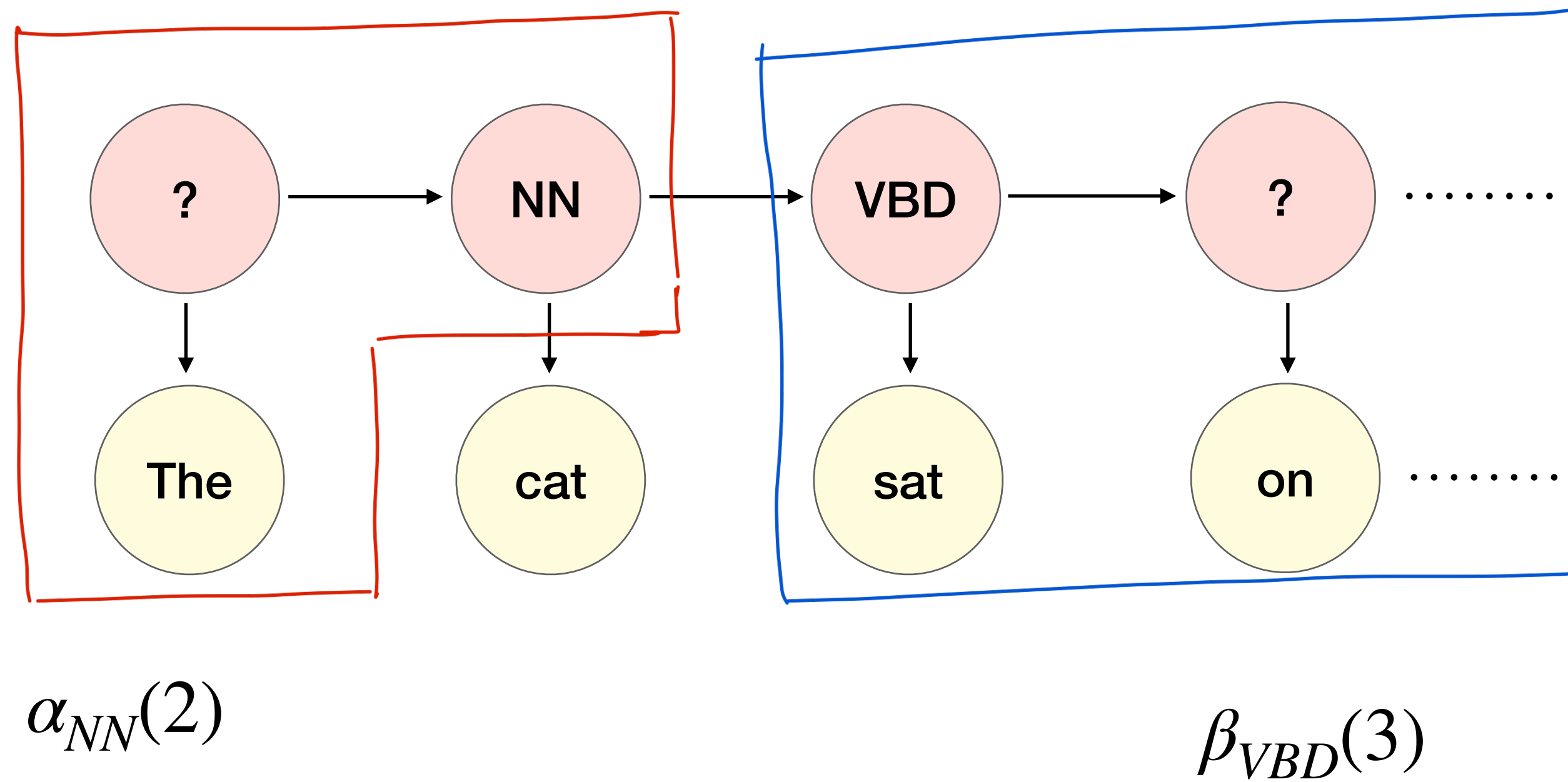
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i.e. the marginal probability of seeing observations x_j, \dots, x_m given $y_j = s$

- Let us now try to express expected counts in terms of α, β

Forward-backward algorithm



$$\alpha_s(j) = P(x_1, \dots, x_{j-1}, y_j = s \mid \theta, \phi)$$

$$\beta_s(j) = P(x_j, \dots, x_m \mid y_j = s, \theta, \phi)$$

Forward-backward algorithm

- Observation likelihood,

$$\begin{aligned} Z = P(x_1, x_2, \dots, x_m | \theta, \phi) &= \sum_s P(x_1, x_2, \dots, x_{j-1}, y_j = s, x_j, \dots, x_m | \theta, \phi) \\ &= \sum_s P(x_1, x_2, \dots, x_{j-1}, y_j = s | \theta, \phi) P(x_j, \dots, x_m | y_j = s, \theta, \phi) \\ &= \sum_s \alpha_s(j) \beta_s(j) \end{aligned}$$

for any $j \in 1, \dots, m$

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$$Z = P(x_1, x_2, \dots, x_m | \theta, \phi) = \sum_s \alpha_s(j) \beta_s(j) \quad \text{for any } j \in 1, \dots, m$$

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- Now, we can compute the following in terms of α, β :

$$P(y_j = s | X, \theta, \phi) = \frac{P(X, y_j = s | \theta, \phi)}{P(X | \theta, \phi)} = \frac{P(x_1, \dots, x_{j-1}, y_j = s | \theta, \phi) P(x_j, \dots, x_m | y_j = s, \theta, \phi)}{Z} = \frac{\alpha_s(j) \beta_s(j)}{Z}$$

Forward-backward algorithm

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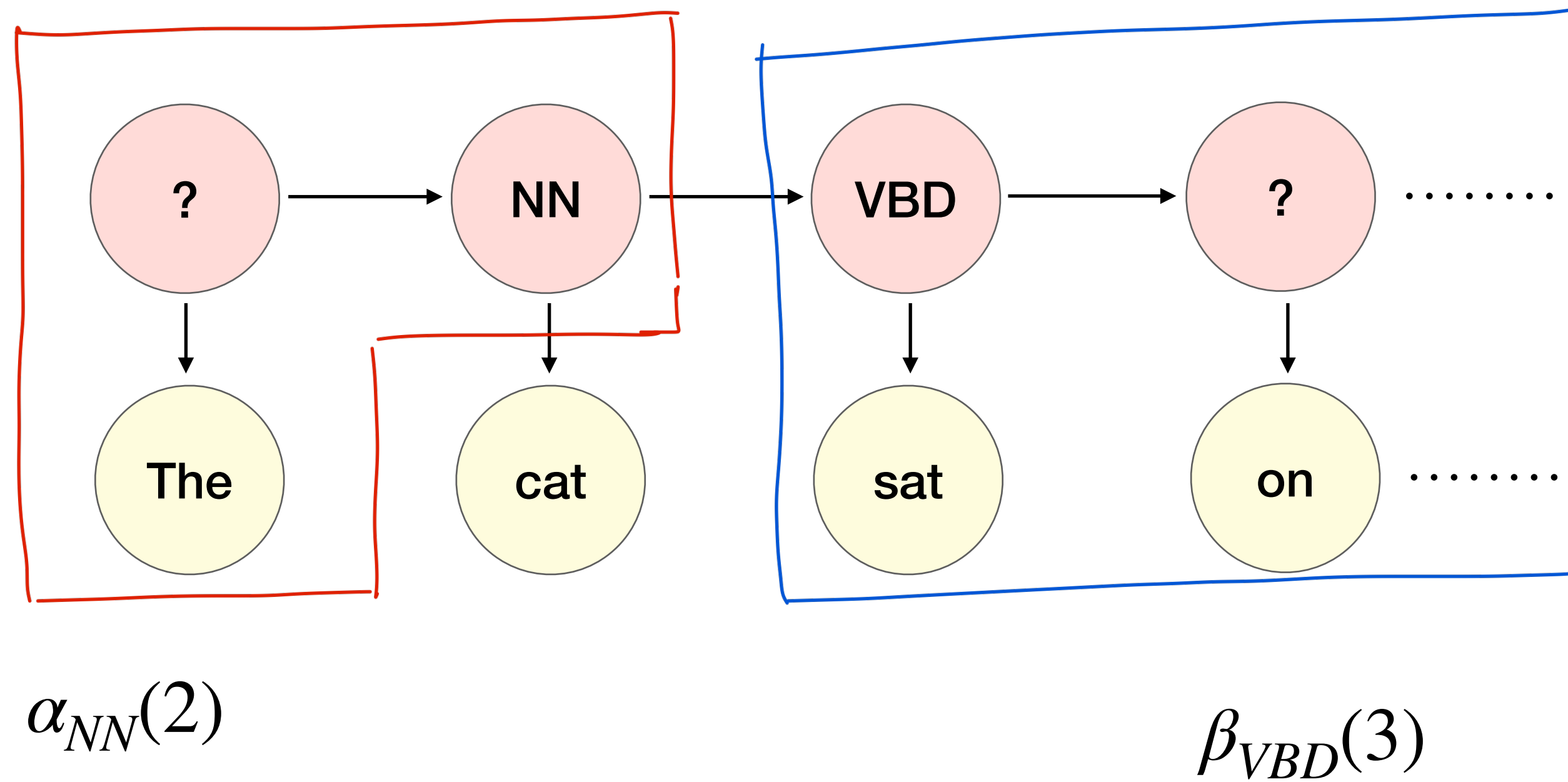
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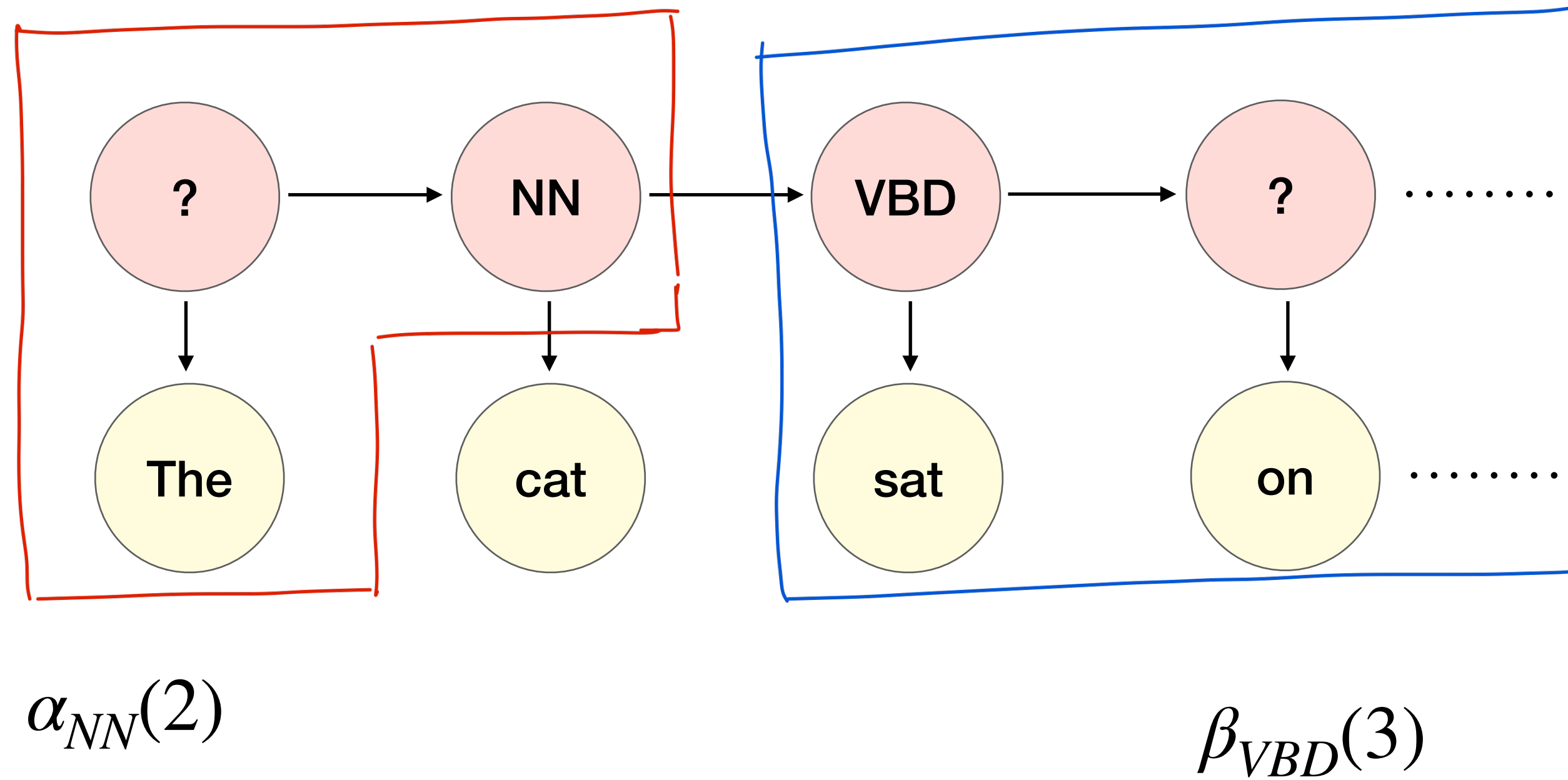
$$P(y_j = s | X, \theta, \phi) = \frac{P(X, y_j = s | \theta, \phi)}{P(X | \theta, \phi)} = \frac{P(x_1, \dots, x_{j-1}, y_j = s | \theta, \phi) P(x_j, \dots, x_m | y_j = s, \theta, \phi)}{Z} = \frac{\alpha_s(j) \beta_s(j)}{Z}$$

- and $P(y_j = s, y_{j+1} = s' | X, \theta, \phi) = \frac{\alpha_s(j) \theta_{s \rightarrow s'} \phi_{s \rightarrow x_j} \beta_{s'}(j+1)}{Z}$

Forward-backward algorithm

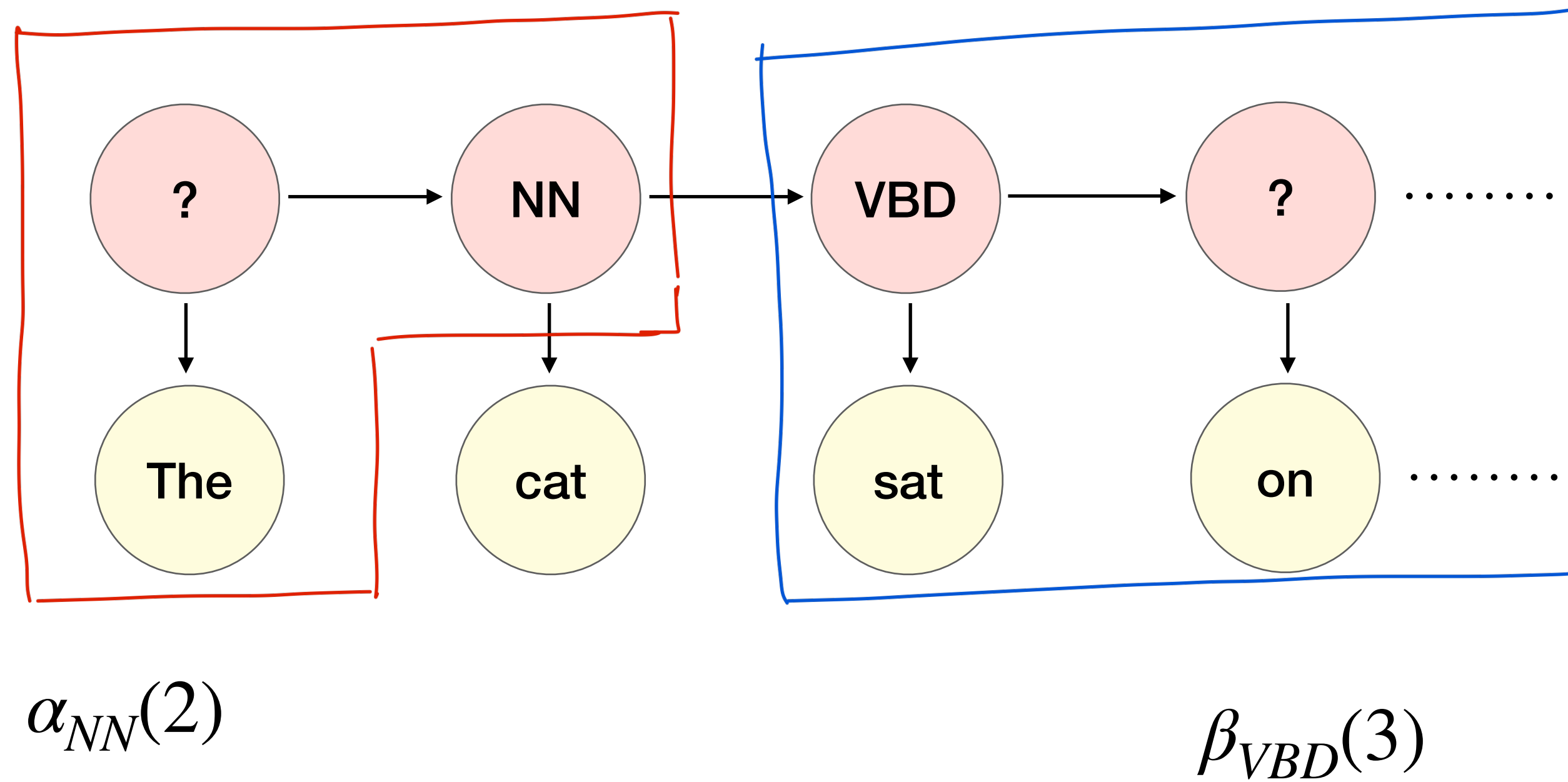


Forward-backward algorithm



- $$P(y_j = NN, y_{j+1} = VBD | X, \theta, \phi) = \frac{\alpha_{NN}(2) \theta_{NN \rightarrow VBD} \phi_{NN \rightarrow cat} \beta_{VBD}(3)}{Z}$$

Forward-backward algorithm



- $$P(y_j = NN, y_{j+1} = VBD | X, \theta, \phi) = \frac{\alpha_{NN}(2) \theta_{NN \rightarrow VBD} \phi_{NN \rightarrow cat} \beta_{VBD}(3)}{Z}$$
- $$P(y_j = NN | X, \theta, \phi) = \frac{\alpha_{NN}(2) \beta_{NN}(2)}{Z}$$

Forward-backward algorithm

- $$P(y_j = s \mid X, \theta, \phi) = \frac{\alpha_s(j)\beta_s(j)}{Z}$$

$$P(y_j = s, y_{j+1} = s' \mid X, \theta, \phi) = \frac{\alpha_s(j) \theta_{s \rightarrow s'} \phi_{s \rightarrow x_j} \beta_{s'}(j+1)}{Z}$$

- Given these, we can now estimate the expected counts:

$$\overline{Count}(s \rightarrow s') = \sum_i \sum_{j=1}^m P(y_j = s, y_{j+1} = s' \mid X_i, \theta, \phi)$$

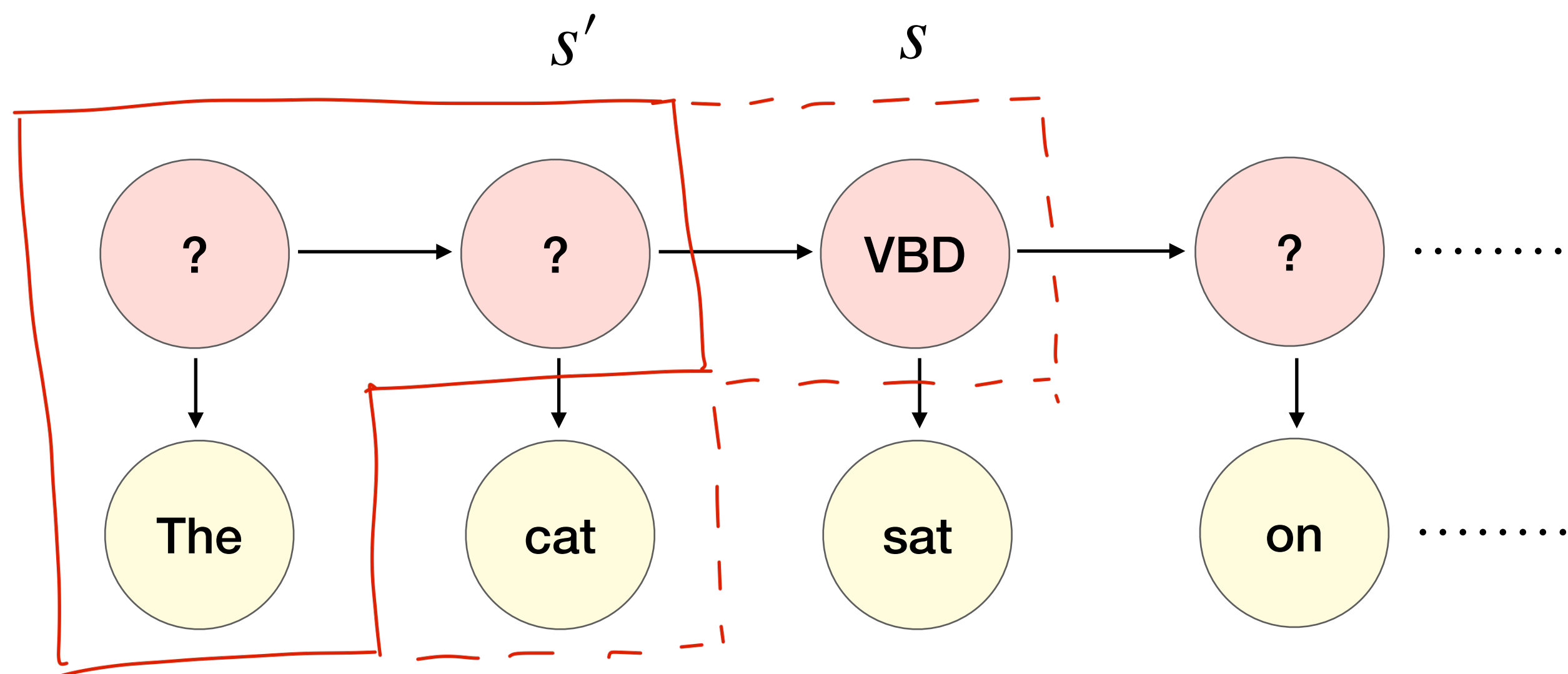
$$\overline{Count}(s \rightarrow o) = \sum_i \sum_{j: X_{ij} = o} P(y_j = s \mid X_i, \theta, \phi)$$

for all s, s', o

Dynamic programming

$$\begin{aligned}
 \alpha_s(j) &= P(y_j = s, x_1, \dots, x_{j-1}) \\
 &= \sum_{s'} P(y_{j-1} = s', x_1, \dots, x_{j-2}) P(x_{j-1} | y_{j-1} = s') P(y_j = s | y_{j-1} = s') \\
 &= \sum_{s'} \alpha_{s'}(j-1) \phi_{s' \rightarrow x_{j-1}} \theta_{s' \rightarrow s}
 \end{aligned}$$

α and β can be computed very efficiently!



Dynamic programming

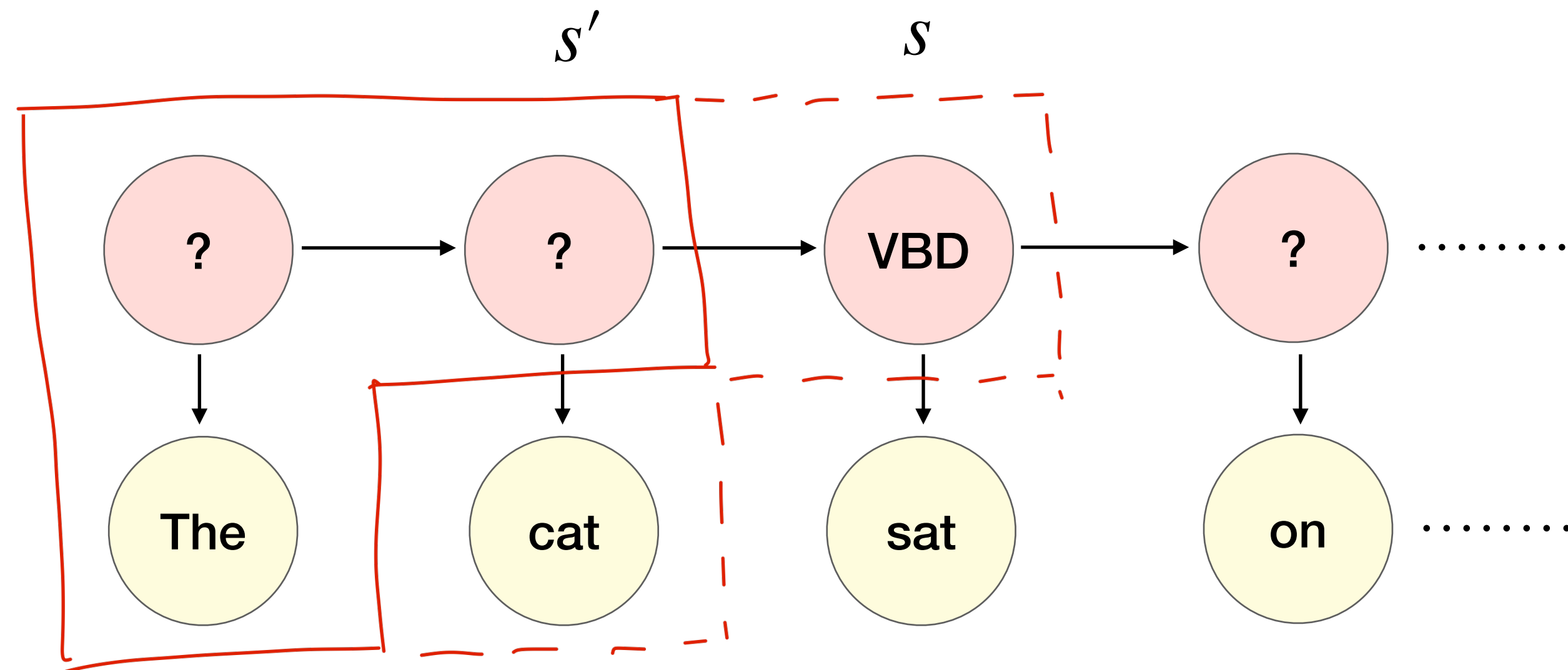
$$\alpha_s(j) = P(y_j = s, x_1, \dots, x_{j-1})$$

$$= \sum_{s'} P(y_{j-1} = s', x_1, \dots, x_{j-2}) P(x_{j-1} | y_{j-1} = s') P(y_j = s | y_{j-1} = s')$$

$$= \sum_{s'} \alpha_{s'}(j-1) \phi_{s' \rightarrow x_{j-1}} \theta_{s' \rightarrow s}$$

$$\alpha_s(1) = \theta_{\emptyset \rightarrow s}$$

α and β can be computed very efficiently!



Dynamic programming



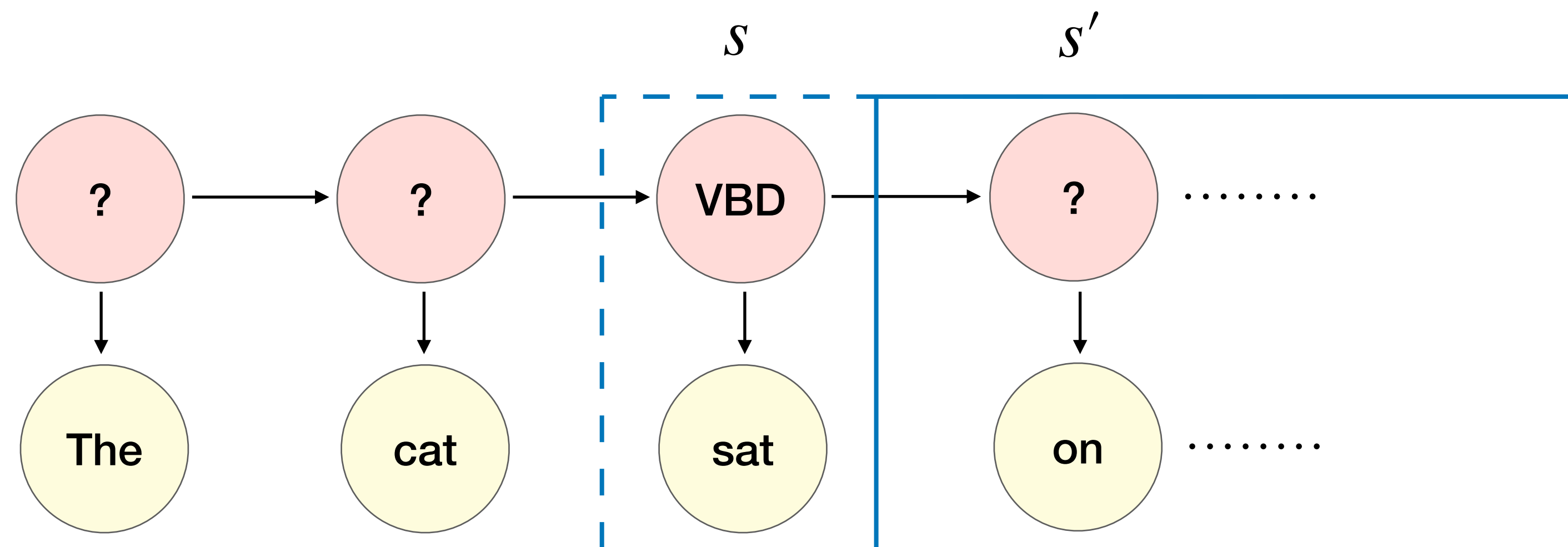
- Similarly,

$$\beta_s(j) = P(x_j, \dots, x_m | y_j = s)$$

$$= \sum_{s'} P(x_{j+1}, \dots, x_m | y_{j+1} = s') P(y_{j+1} = s' | y_j = s) P(x_j | y_j = s)$$

$$= \phi_{s \rightarrow x_j} \sum_{s'} \beta_{s'}(j+1) \theta_{s \rightarrow s'}$$

α and β can be
computed very
efficiently!



Dynamic programming



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$$\beta_s(j) = P(x_j, \dots, x_m | y_j = s)$$

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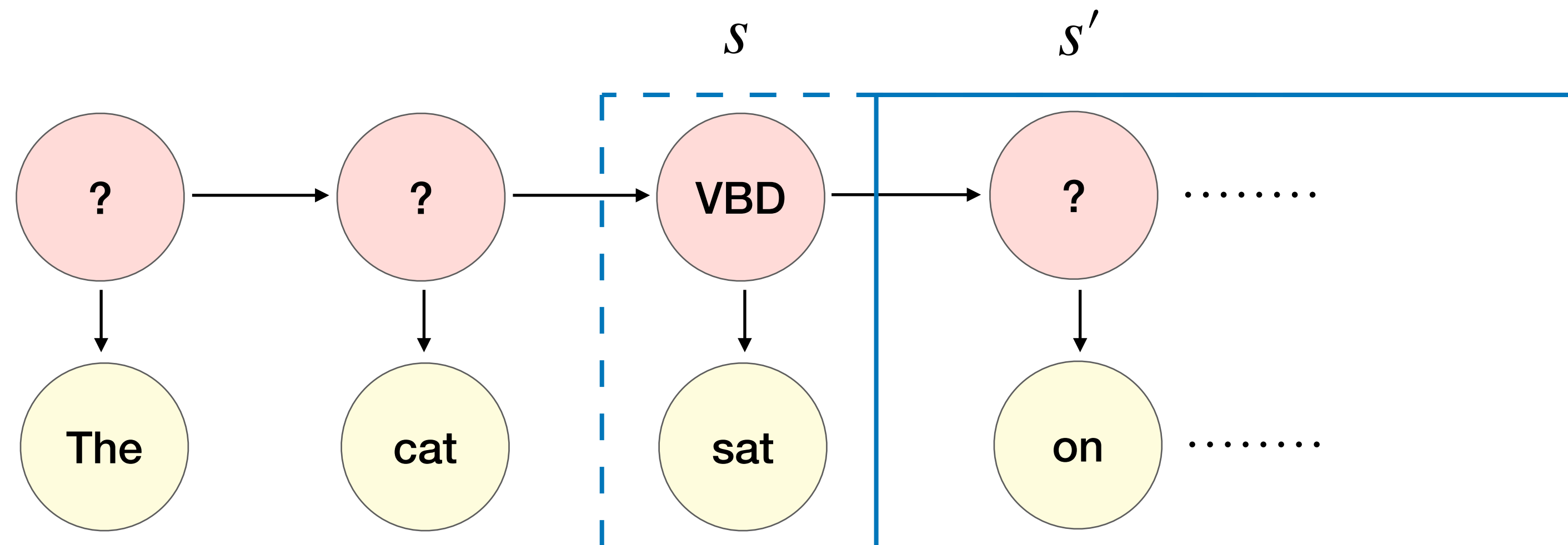
What is the base case?

A) $\beta_s(m) = \phi_{s \rightarrow x_m}$

B) $\beta_s(m) = 1$

C) $\beta_s(m) = \theta_{\emptyset \rightarrow s}$

α and β can be computed very efficiently!



Dynamic programming



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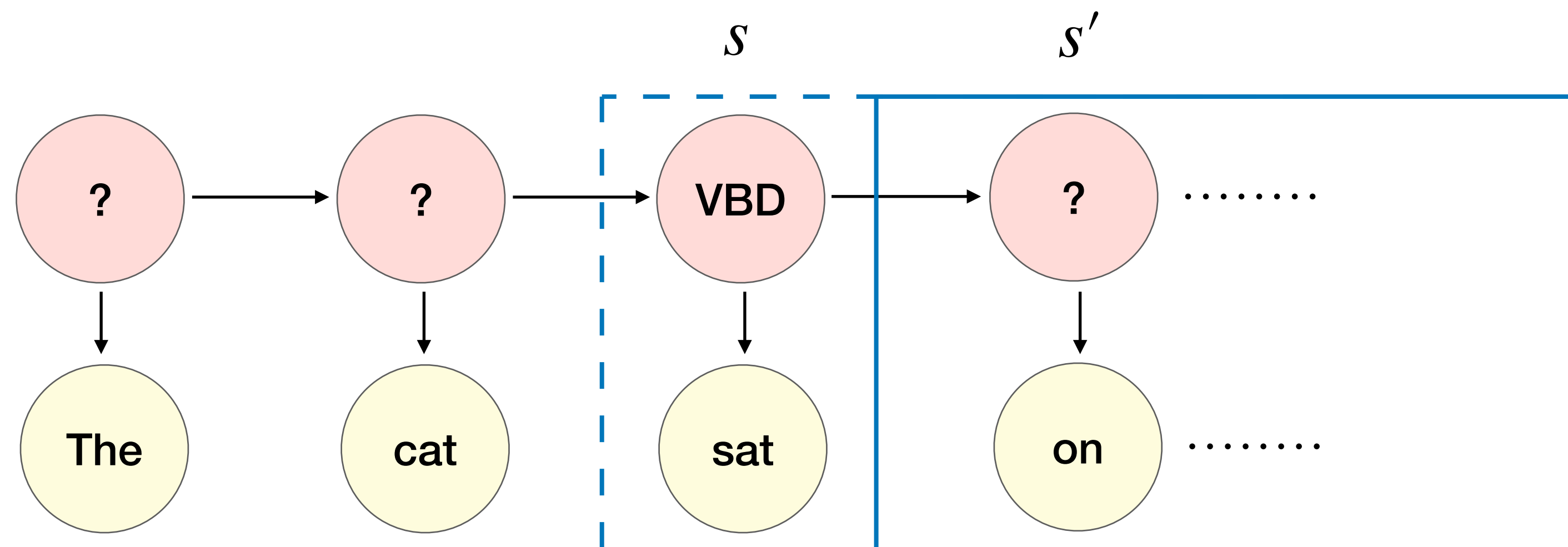
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Dynamic programming



- $\alpha_s(j) = \sum_{s'} \alpha_{s'}(j-1) \phi_{s' \rightarrow x_{j-1}} \theta_{s' \rightarrow s}$
- $\beta_s(j) = \phi_{s \rightarrow x_j} \sum_{s'} \beta_{s'}(j+1) \theta_{s \rightarrow s'}$
- Compute for all $s \in S, j \in [1, m]$

Dynamic programming



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What is the runtime of this dynamic programming algorithm?

- A) $O(|S| \cdot m)$
- B) $O(|S| \cdot m^2)$
- C) $O(|S|^2 \cdot m)$

Dynamic programming



- $\alpha_s(j) = \sum_{s'} \alpha_{s'}(j-1) \phi_{s' \rightarrow x_{j-1}} \theta_{s' \rightarrow s}$
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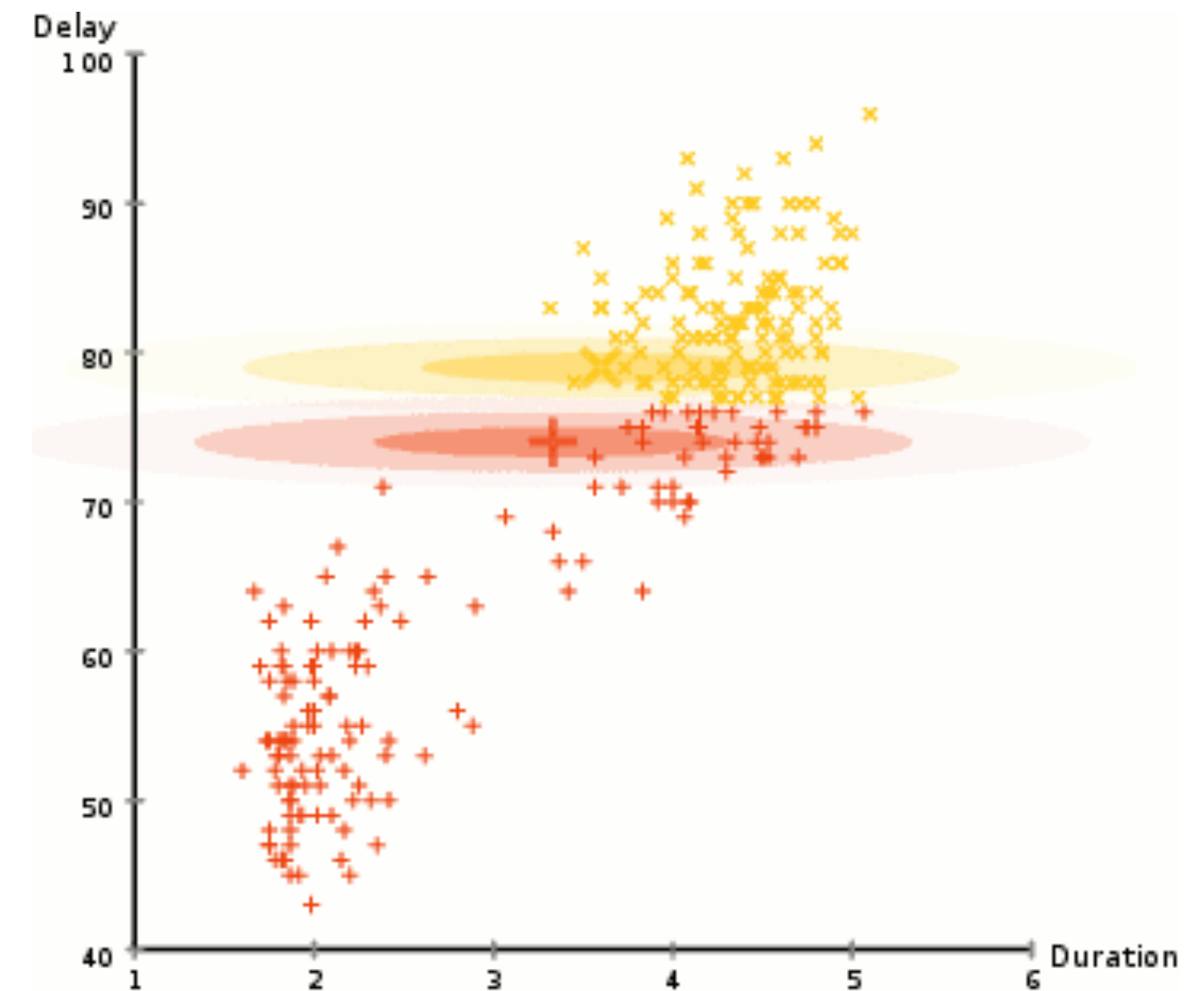
A) $O(|S| \cdot m)$

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EM applications

- Any task with unobserved latent variables
- In NLP:
 - Sequence modeling
 - Syntactic parsing (inside-outside algorithm)
- Clustering (cluster IDs = hidden variables)
- Computer vision (segmentation, activity recognition)
- Quantitative genetics, psychometrics, medical image reconstruction, structural engineering ...

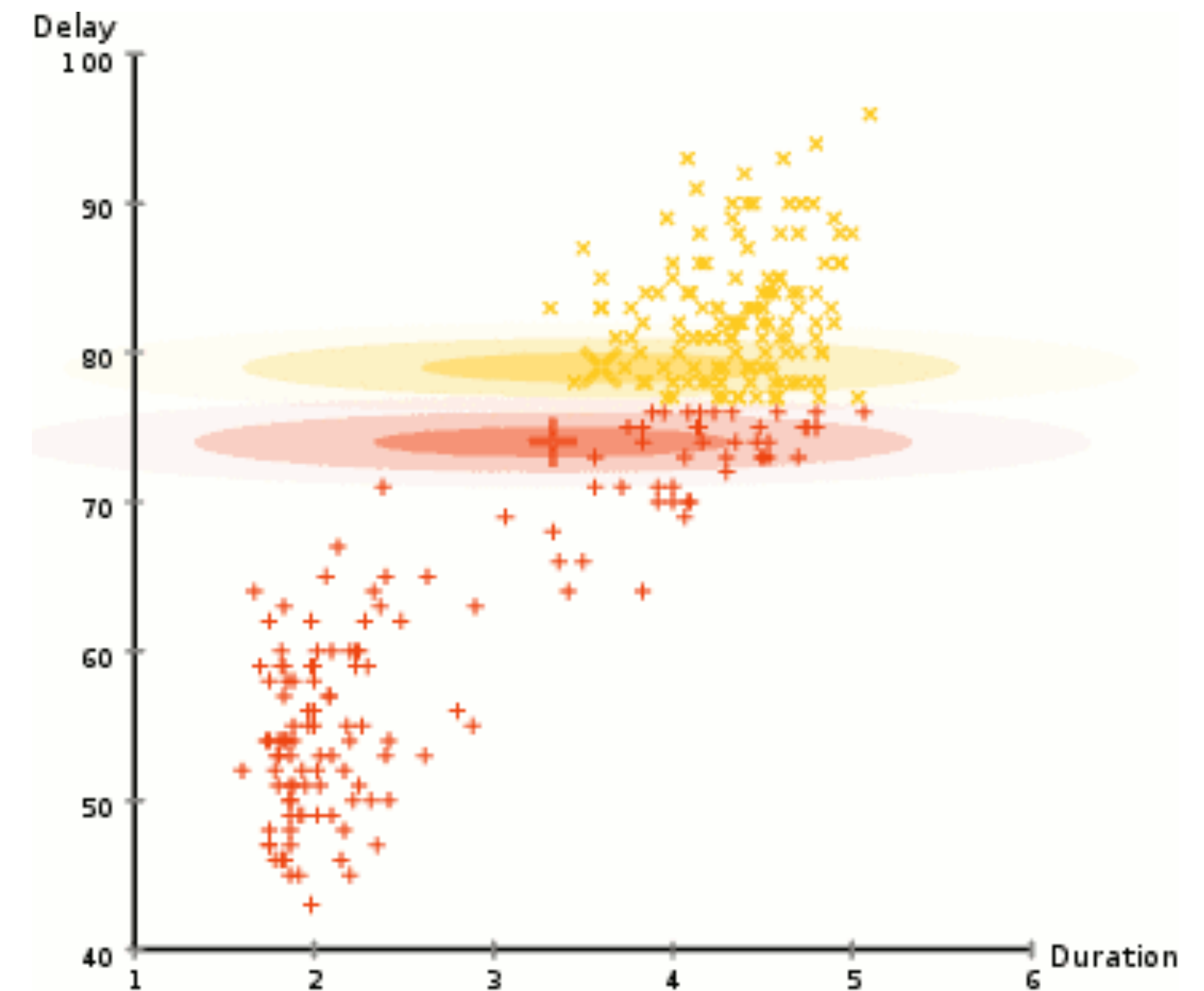


Clustering

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